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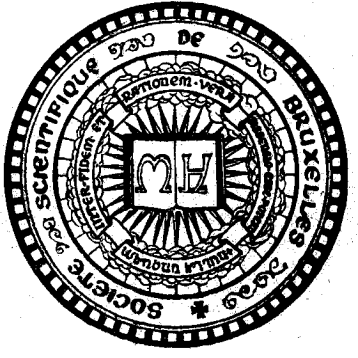
TOME SOIXANTE-TREIZIÈME

SÉRIE I

SCIENCES MATHÉMATIQUES ASTRONOMIQUES ET PHYSIQUES

TROISIÈME ET DERNIER FASCICULE

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Deuxième Section

SCIENCES PHYSIQUES ET CHIMIQUES

On the statistical laws of linguistic distributions

BY

V. BELEVITCH (*)

SUMMARY

The rank-frequency diagrams of statistical linguistics are reinterpreted as distribution curves of the cumulative probability of types in the catalogue versus the probability of tokens in the text. For such distributions, the closure condition $\sum p_i = 1$ (which does not hold in general statistics for the independent variable) imposes certain relations between the mean, the variance, the number of elements in the catalogue and the average information content (negative entropy). Sections 2 to 4 are devoted to the mathematics of these relations, especially to their particular forms for truncated normal distributions, first and second order Taylor approximations to an arbitrary distribution law (take the form of Zipf's and Mandelbrot's laws respectively. Experimental data lead to accept the truncated normal distribution with $\sigma \cong 2,8$ bits as the general law for words. Data on letter and phoneme distributions seem to indicate that the standard deviation has the universal value $\sigma \cong 1,4$ bits.

RANK-FREQUENCY DIAGRAMS

It is an experimental fact that by counting frequencies of occurrence of various elements (letters, phonemes, words) in homogeneous texts written in a given language, and dividing them by the total number of elements of the same nature in the text, one often obtains relative frequencies that are stable (independent of the length of the text for sufficiently long texts). These relative frequencies define the a priori probabilities of the elements. Having used texts to obtain probabilities of various linguistic elements, one constructs catalogues of elements of identical nature (ie. alphabet for letters, lexicon for words, etc.) in which each

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In normal statistical practice one defines discrete distributions by specifying the number N_i of elements having a dimension or some other measurable characteristic x_i . If $N = \sum N_i$ is the total number of elements, the ratio $f_i = N_i/N$ is the relative frequency, or probability of finding the value x_i for the dimension x . The cumulative probability is defined by

$$\phi_i = \sum_{j=1}^i f_j \tag{1}$$

and distribution curves are obtained by plotting the cumulative probability ϕ_i versus the dimension x_i .

Most linguistic elements (i.e. letters) have no measurable characteristic (dimension) according to which they could be ordered, except their text probabilities themselves. As there is no point in defining a distribution curve by a text probability versus a text probability, the only alternative is to plot the probability in the catalogue f_i versus the probability in the text p_i . The probability in the catalogue is defined by reference to experiments with an urn containing once each element i , marked with its text probability p_i : the catalogue probability f_i is the probability of drawing from the urn the text probability p_i . For a catalogue of N elements, N_i of which have the text probability p_i , f_i is the ratio N_i/N , and the cumulative probability (1) in the catalogue is

$$\phi_i = \frac{1}{N} \sum_{j=1}^i N_j \tag{2}$$

But the sum $\sum N_i$ in (2), i.e. the number of elements of text probabilities $\geq p_i$, is precisely the rank i in the catalogue, since elements are ranged in order of non-increasing frequencies. As a consequence (2) becomes $\phi_i = i/N$ and gives the relative rank in the catalogue. The distribution curves thus defined only differ from the usual rank-frequency diagrams, where the rank of each element in the catalogue is plotted versus its probability in the text, by the factor $1/N$ transforming absolute rank into relative rank. As a conclusion, rank-frequency diagrams can be interpreted as ordinary distribution curves giving the cumulative probability in the catalogue versus the probability in the text. This establishes a relation between the paradigmatic and syntagmatic aspects of the language.

Information theory attributes to each element of probability p_i an information measure (negative entropy) $x_i = -\log p_i$, and it is therefore convenient to use logarithmic scales for text probabilities. When a logarithm of base 2 is used in the above definition, the information is measured in *bits*. In the analytical expressions, it is more convenient to use natural logarithms; this is equivalent to adopt $\log_2 e = 1.44$ bits as natural unit of information (*bin*).

In general statistics, the range of the independent variable x is unrestricted. In statistical linguistics, x is related to a probability by and is essentially positive. An additional restriction results from the *closure condition*

$$(4) \quad \sum_N N_i p_i = 1$$

which is transformed into

$$(5) \quad \sum N_i e^{-x_i} = 1$$

or, since the probability in the catalogue was defined as $f_i = N_i/N$, and

$$(6) \quad \sum \frac{N_i}{N} e^{-x_i} = 1$$

As a conclusion, the *closure condition determines the absolute number of elements in the catalogue from their relative distribution law*

2. MEAN VALUES

When dealing with mean values, one should clearly distinguish between *averages over the text* and *averages over the catalogue*. The mean m and the variance σ^2 as defined by

$$(7) \quad m = \frac{\sum N_i x_i}{\sum N_i}; \sigma^2 = \frac{\sum N_i x_i^2 - m^2}{\sum N_i - m}$$

are averages over the catalogue. In the text, the number N_i of elements of measure x_i must be weighted proportionally to their probability of occurrence p_i , and the average information is

$$h = \frac{\sum N_i p_i x_i}{\sum N_i p_i}$$

$$(8) \quad h = - \sum N_i p_i \log p_i$$

in the sense of information theory. Finally (8) becomes

$$(9) \quad \frac{h}{N} = \sum f_i x_i e^{-x_i}$$

We now consider the case where the range of x , for which f takes significant values, is sufficiently small, i.e. *narrow distributions*, so that the function e^{-x} can be approximated by the Taylor expansion

$$(10) \quad e^{-x} = e^{-m} \left[1 - (x-m) + \frac{1}{2} (x-m)^2 \right]$$

around the mean m . Condition (6), combined with (7), yields

$$(11) \quad e^{-m} (1 + \sigma^2/2) = 1/N$$

Similarly, in (9), the expansion of $x e^{-x}$ is deduced from (10) by writing $x = (x-m) + m$ and neglecting the third order term. One finds

$$x e^{-x} = m e^{-m} \left[1 - (x-m) + \frac{m}{1} (x-m) + \frac{1}{2} (x-m)^2 \right]$$

and (9) becomes

$$(12) \quad \frac{h}{N} = m e^{-m} \left[1 + \frac{1}{2} \left(\frac{m}{1} - \frac{m}{1} \right) \sigma^2 \right]$$

Since σ^2 is small for a narrow distribution, (11) becomes approximately

$$(13) \quad m = \log N + \frac{1}{2} \sigma^2$$

On the other hand, the ratio of (12) and (11) gives, with the same approximation,

$$(14) \quad h = \log N - \frac{1}{2} \sigma^2$$

or

$$(15) \quad h = m - \sigma^2$$

If all N elements have the same text probability $1/N$, the information is the constant $\log N$. By comparison, formulae (13) and (14) show that the *effect of a small spread in the probabilities is to reduce the average information in the text* (and this is well known from information theory) *and to increase by the same amount the average over the catalogue*. Furthermore, the variance of the distribution is precisely the difference between both averages.

It is often convenient to approximate discrete distributions with a large number of elements by continuous distributions. The number of elements of dimension comprised between x and $x + dx$ is $N d\phi(x) = N \phi'(x) dx = N f(x) dx$, where $\phi(x)$ is the distribution function and $f(x) = \phi'(x)$ the probability density. A Gaussian, or normal, distribution of mean m and variance σ^2 is defined by the probability density

$$f(x) = \frac{1}{\sigma\sqrt{2\pi}} e^{-\frac{x-m}{\sigma}} \quad (16)$$

The range of the normal distribution is $-\infty \leq x \leq \infty$ and, since x is essentially positive in linguistic applications, the distributions cannot be rigorously normal. In the following we will consider truncated normal distributions where (16) is restricted to some positive interval $x_0 \leq x \leq x_b$, the density assuming the value 0 outside. We will start, however, by the case where the truncation is made at points sufficiently far away from the mean, so that the tails of the distribution are negligible anyway.

For a continuous distribution, conditions (6) and (9) are replaced by

$$\frac{1}{h} = \int_{x_0}^{x_b} f(x) e^{-x} dx; \quad \frac{N}{h} = \int_{x_0}^{x_b} x e^{-x} f(x) dx \quad (17)$$

In these expressions also, we first replace the integration limits by $\pm \infty$, neglecting the truncation. Absolute convergence is still ensured, for the increase of $|x|$ and e^{-x} for $x = -\infty$ is less rapid than the decrease of (16).

For the density (16), the integrals (17) are reduced to the error integral by transforming the exponent according to

$$x = \frac{2\sigma^2}{\sigma^2} \frac{(x-m)^2}{2\sigma^2} - m \quad (18)$$

By the linear transformation

$$z = \sigma + \frac{\sigma}{m-x} \quad (19)$$

the first equation (17) becomes

$$\frac{1}{h} = \frac{N}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{-z^2/2} dz = e^{-m-\sigma^2/2} \quad (20)$$

and is equivalent to (13). The second equation (18) becomes

$$\frac{N}{h} = \frac{N}{\sigma\sqrt{2\pi}} \int_{-\infty}^{\infty} (m - \sigma^2 z) e^{-z^2/2} dz \quad (21)$$

The term in z in the integrand is an odd function and does not contribute to the integral, so that (21) reduces to

$$\frac{N}{h} = (m - \sigma^2) e^{-m-\sigma^2/2} \quad (22)$$

and, by comparison with (20), one obtains (15).

The remaining part of this section is devoted to the derivation of the rigorous formulae replacing (20) and (22) and taking the truncation into account. For the error integral, we will use the notation

$$\Phi(n) = \frac{1}{\sqrt{2\pi}} \int_n^{\infty} e^{-u^2/2} du \quad (23)$$

so that the distribution function corresponding to the density (16), with neglected truncation, is

$$\phi(x) = \Phi\left(\frac{\sigma}{m-x}\right) \quad (24)$$

If truncation is taken into account, the density cannot be (16), but must be corrected by a factor $a > 1$ in order to normalize to unity the integral in the finite range. Similarly the distribution function is no longer (24) but is of the form

$$\phi(x) = a \Phi\left(\frac{\sigma}{m-x}\right) - b \quad (25)$$

deduced from the density (16) multiplied by a , with an integration constant denoted $-b$. These constants are determined by the conditions $\phi(x_b) = 0$, $\phi(x_0) = 1$, and one obtains

$$a = \frac{\Phi\left(\frac{\sigma}{m-x_0}\right)}{\Phi\left(\frac{\sigma}{m-x_b}\right)}; \quad b = \frac{\Phi\left(\frac{\sigma}{m-x_0}\right) \Phi\left(\frac{\sigma}{m-x_b}\right)}{\Phi\left(\frac{\sigma}{m-x_0}\right) - \Phi\left(\frac{\sigma}{m-x_b}\right)} \quad (26)$$

For the truncated distribution, the integrals (11), applied to the density $a f(x)$, become

$$(27) \quad 1/N = a e^{-m\sigma^2/2}$$

$$(28) \quad h = \frac{1}{\sigma} \left[c(m - \sigma^2) + \frac{\sqrt{2\pi}}{\sigma} (e^{-z^2/2} - e^{-z^2/2}) \right]$$

where z_0 and z_1 have the values resulting from (19) with $x = x_0$ or x_1 , and where c denotes the error integral with the limits z_0 and z_1 , thus

$$(29) \quad c = \Phi \left(\sigma + \frac{x_0 - m}{\sigma} \right) - \Phi \left(\sigma + \frac{x_1 - m}{\sigma} \right)$$

In (25), the term b introduces a correction at high probabilities, which is negligible in most applications because $x_n - m$ is negative and equals several times σ . On the contrary, the effect of the truncation at low frequencies is often not negligible, for statistics do not generally extend sufficiently far above the mean. The practical form of (27) is thus deduced from the approximation $x_n = -z_1$ and is

$$(30) \quad N = e^{-m\sigma^2/2} \frac{\Phi \left(\sigma + \frac{x_0 - m}{\sigma} \right)}{\Phi \left(\sigma + \frac{x_1 - m}{\sigma} \right)}$$

In a complete statistical count, the lowest frequency corresponds to a single occurrence in the text, thus to probability $p_0 = 1/L$, where L is the length of the text. By (3), one has

$$(31) \quad x_0 = \log L$$

and (30) gives a relation between the length of the text and the extension of the vocabulary. It is obvious that (30) reduces to (13) for $x_1 = \infty$.

4. APPROXIMATIONS TO NORMAL DISTRIBUTIONS

We consider first an arbitrary distribution function $\varphi(x)$ in the neighbourhood of a point x_0 . The Taylor expansion is

$$\varphi(x) = \varphi(x_0) + (x - x_0)f(x_0)$$

where $f(x)$ is the corresponding probability density. The logarithmic slope of distribution function is approximately

$$(32) \quad \log \frac{\varphi(x)}{\varphi(x_0)} = \log \left[1 + \frac{f(x_0)}{\varphi(x_0)} (x - x_0) \right] \approx (x - x_0) \frac{f(x_0)}{\varphi(x_0)}$$

If one considers an element $x = x_i$ corresponding to a text probability $p_i = e^{-x_i}$, its rank i is given by $N\varphi(x_i)$. Introducing the similar notions p_0 and i_0 for the reference point x_0 , (32) becomes

$$(33) \quad \log \frac{p_i}{p_0} = -A \log \frac{i}{i_0}$$

with

$$(34) \quad A = \frac{f(x_0)}{\varphi(x_0)}$$

Equation (33) is independent from any assumption on the distribution law, and merely shows that (34) measures the slope of the tangent, at x_0 , to the rank-frequency characteristic with logarithmic scales for both coordinates.

Expression (33), or

$$(35) \quad \frac{i}{i_0} = \left(\frac{p_0}{p_i} \right)^{-1/A}$$

is similar to Zipf's law, but with a variable slope. By taking into account second order terms in the Taylor expansion, it is possible to obtain a correction similar to the one introduced in Zipf's law by Mandelbrot, i.e. to arrive at a form

$$(36) \quad \frac{i}{i_0} = s \left(\frac{p_0}{p_i} \right)^{-1/B} - t$$

instead of (35), or, equivalently, to

$$(37) \quad p_i = P(i + \rho)^{-B}$$

with

$$(38) \quad P = p_0(i_0 s)^B; \rho = i_0 t$$

The best values of the parameters are obtained by identifying the second-order Taylor expansion of the right-hand side of (36), i.e.

$$s e^{(x_i - x_0)/B} - t = s - t + s \frac{B}{x_i - x_0} + \frac{2B^2}{2(x_i - x_0)^2}$$

(1) The possibility of deducing Mandelbrot's law with $B \leq 1$ from the asymptotic expansion of the error integral was mentioned to the writer by A. OETTINGER, and originated the present investigation.

When the approximating point x_0 is the point of highest frequency v_m , (44) gives $B > 1$, since the average information h is certainly larger than the minimum information x_m . When x_0 increases

$$B = \frac{\sigma^2}{\sigma^2 + h - x_0} \quad (44)$$

By (15), (41) becomes

$$s = 1 + \frac{(m-x_0)^2}{\sigma^2}; t = \frac{(m-x_0)^2}{\sigma^2} \quad (43)$$

of the Zipf approximation. With two terms in (42), (39) gives valid for large positive values of u . When the first term alone of (42) is considered, the value (41) is obtained for the exponent (34)

$$\phi(-n) = \frac{e^{-n^2/2}}{1 - \frac{n}{n^2} + \dots} \quad (42)$$

use the asymptotic expansion (1) statistics are the most reliable. We will thus assume $x_0 \ll m$ and for the behavior of the characteristics at high frequencies, where the for the linguistic applications it is of special importance to discuss some approximations are introduced, and truncation neglected. Simple expressions for the other parameters are only obtained if

$$B = \frac{\sigma^2}{m - x_0} \quad (41)$$

For a normal distribution (truncated or not), (40) becomes at x_0 is positive:

and one has $t > 0$ as long as the curvature of the characteristic

$$B = \frac{f'(x_0)}{f(x_0)} \quad (40)$$

$$s = \frac{[f(x_0)]^2}{\varphi(x_0)f'(x_0)}; t = s - 1 \quad (39)$$

This gives

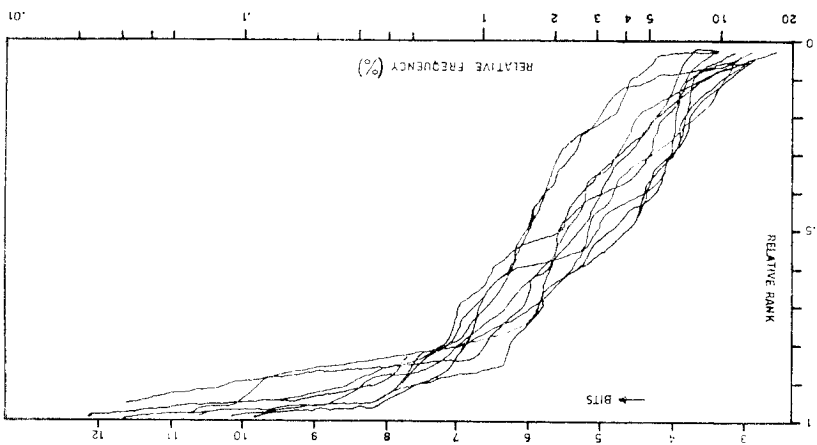
$$\frac{\varphi(x_i)}{\varphi(x_0)} = 1 + \frac{\varphi(x_0)}{(x_i - x_0)f(x_0)} + \frac{2\varphi(x_0)}{(x_i - x_0)^2 f'(x_0)}$$

with the corresponding expansion of the distribution function

(?) Most of these have already been discussed in V. BELEVITCH « Théorie de l'information et statistique linguistique », *Bull. Acad. Roy. Belg. (Cl. des Sc.)* avr. 1956 pp. 419-436.

the upper scale is linear in bits. It appears immediately that all curves are very similar, and all ranges extend from approximately 2.5 to some 11 bits. All distributions are relatively narrow, and it is therefore expected by (13) that the horizontal shift between the various curves is correlated with the size of the alphabet (varying from $N = 21$ to 49 in the examples considered). This is indeed the case, as it appears from Fig. 2 where each curve has been shifted by $\log N$; in other words, the abscissa in fig. 2 is the relative frequency p/p_m with respect to the equiprobable value $p_m = 1/N$.

Fig. 1



Because of the small size of the alphabets, such distributions are relatively irregular, but definite systematic trends can, however, be noticed. On the other hand, thanks to the small size of the alphabets, the statistics are more reliable, and the distributions are well known over their entire ranges. Fig. 1 shows data from a number of languages (2); the lower scale is logarithmic in p , and

5. LETTER AND PHONEME DISTRIBUTIONS

$x_0 = h$.

starting from x_m , B increases and passes through the value 1 for

Fig. 2 clearly shows that all distributions are practically identical and, in particular, have the same variance. Our best estimate of the common value is $\sigma = 1.4$ bits. The corresponding normal distribution shifted by $\sigma^2/2$ relatively to the point of abscissa $\log N$ in accordance with (13), is the dotted curve of fig. 2.

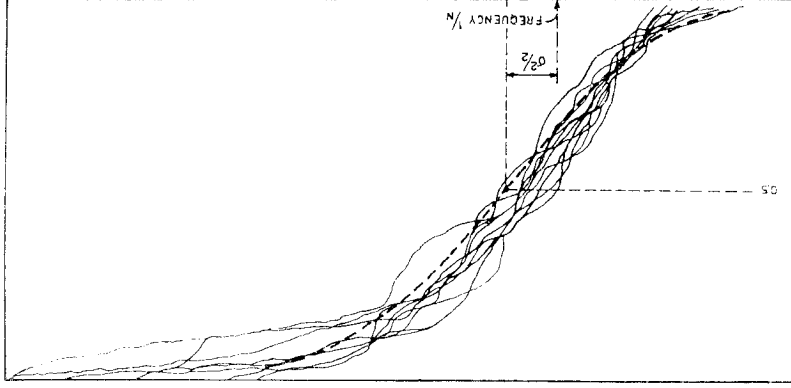


Fig. 2

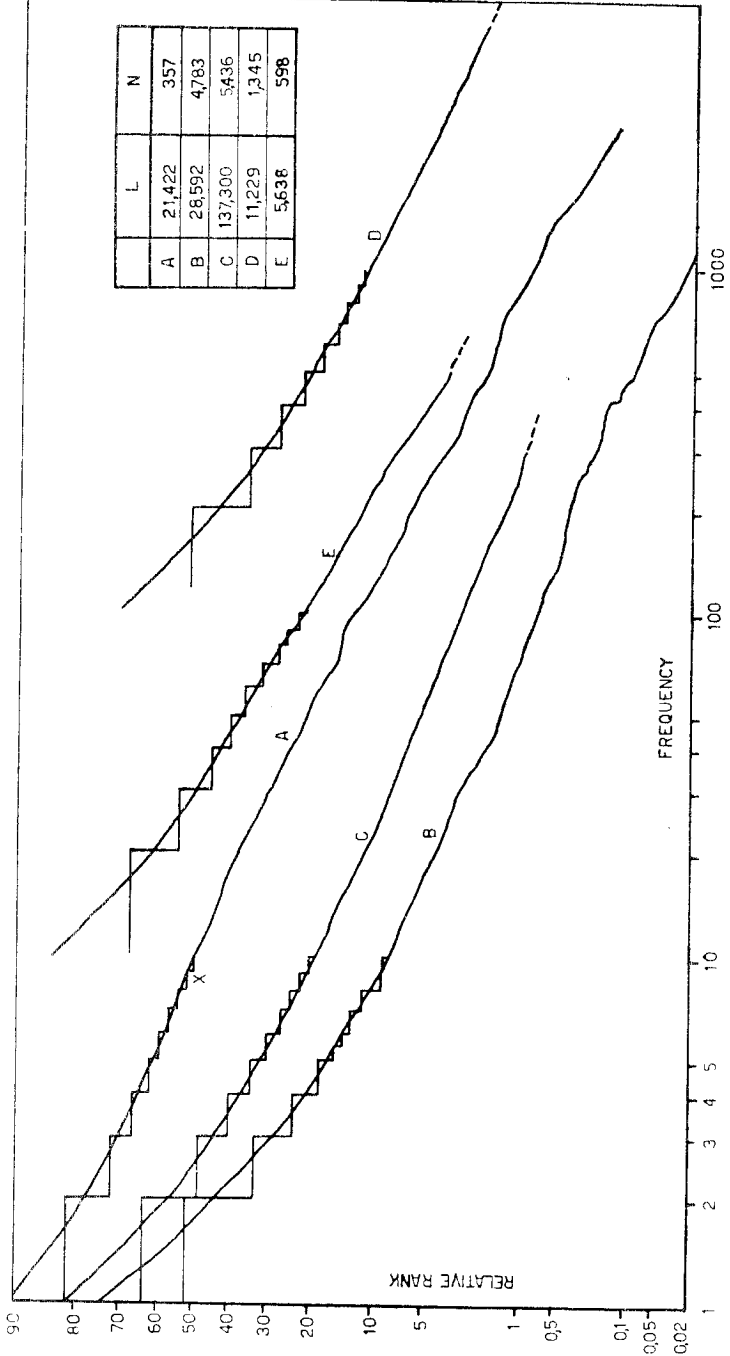
Formulae (13-15) can be checked on the example of Russian phonemes for which particularly detailed data are available ⁽³⁾. The value of $\log N$ is $\log_2 42 = 5.42$ bits, and the published value of h is 4.78 bits. From (15), one obtains $\sigma = 1.33$ bits, and from (13), $m = 6.1$ bits; the last value agrees with the median of the distribution curve deduced from the published data.

6. WORD DISTRIBUTIONS

According to Guiraud ⁽⁴⁾, Zipf's formula, and even Mandelbrot's correction, do not agree with most experimental distributions at low frequencies. The truncated lognormal character of the actual distributions seems to have been suspected by several authors ⁽⁵⁾. If truncation is neglected, a lognormal distribution

⁽³⁾ E.C. CHERRY, M. HALLE, R. JAKOBSON, «Toward the logical description of Languages in their phonemic aspect», *Language*, vol. 29, n° 1, p. 34, March 1953.
⁽⁴⁾ P. GUIRAUD, *Les caractères statistiques du vocabulaire*, Paris, Press. Univ. 1954.
⁽⁵⁾ See F. L. J. ATCHISON, J.A.C. BROWN, *The lognormal distribution*, Cambridge 1957; p. 101.

Fig. 3



on the example of curve B of fig. 3. The corresponding bilogarithmic representation of fig. 5 shows that the value of the initial slope

becomes a straight line on logarithmic probability paper; this is checked for a few examples on Fig. 3, and it will be established herebelow that the deviations from linearity at low frequencies are quantitatively explained by the truncation effect.

The abscissae in fig. 3 are the absolute frequencies, as published in various sources (⁶), but curve E has been shifted by a factor 10 and curve D by a factor 100, to avoid overlapping. The table included in fig. 3 gives the extensions of the vocabulary (N) and of the text (L) mentioned in the source material. The discrete steps at the low frequency ends of the curves arise because the numbers of occurrences of the rarest words are necessarily small integers. Strictly,

the usual definition of the rank-frequency relation would yield a continuous curve passing through the top points of the ladder, but the similar relation based on the complementary rank gives a curve passing through the bottom points. A unique smoothed continuous distribution curve can therefore only be defined by joining the vertical mid-points of the ladder, and this has been done in fig. 3. In particular, the first point of the smoothed distribution, corresponding to frequency 1, is $1 - N_1/2N$, where N_1 is the number of hapaxlegomena and N the total number of different words in the sample.

For curve A, the truncation effect is practically negligible because the statistics extends well above the mean. The slope of the straight line (abscissa interval corresponding to a decrease of the ordinate from 50 to 16%) defines the standard deviation as $\sigma = 2.8$ bits. From this value, and the value of N mentioned in the table, one can compute m by (13), and the corresponding median frequency is $L e^{-m}$; the value thus obtained is shown by a cross on fig. 3. Fig. 4 shows the curve of fig. 3A in bilogarithmic coordinates, and the dotted straight lines are Zipf's approximations at $x = x_n$ and $x = h$, the values of the slope being computed by (44), and the value of h by (15).

(⁶) Curve A is based on data for English function words (G.A. MILLER, E.B. NEWMAN, F.A. FRIEDMAN, « Length frequency statistics for written English », *Inform. and Control*, vol. 1 n° 4 pp. 370-389; Dec. 1958). Curve B corresponds to Russian words in the Captain's Daughter by Pushkin (H.H. JOSSASON, *The Russian word count*, Detroit, 1953). Curves C (New testament and D (St Mark) are based on R. MORGENTHAUER, *Statistik des neuteamentlichen Wortschatzes*, Zurich 1958. Curve E is for French adjectives from all 8 tragedies of Racine (ref. note 4 p. 31).

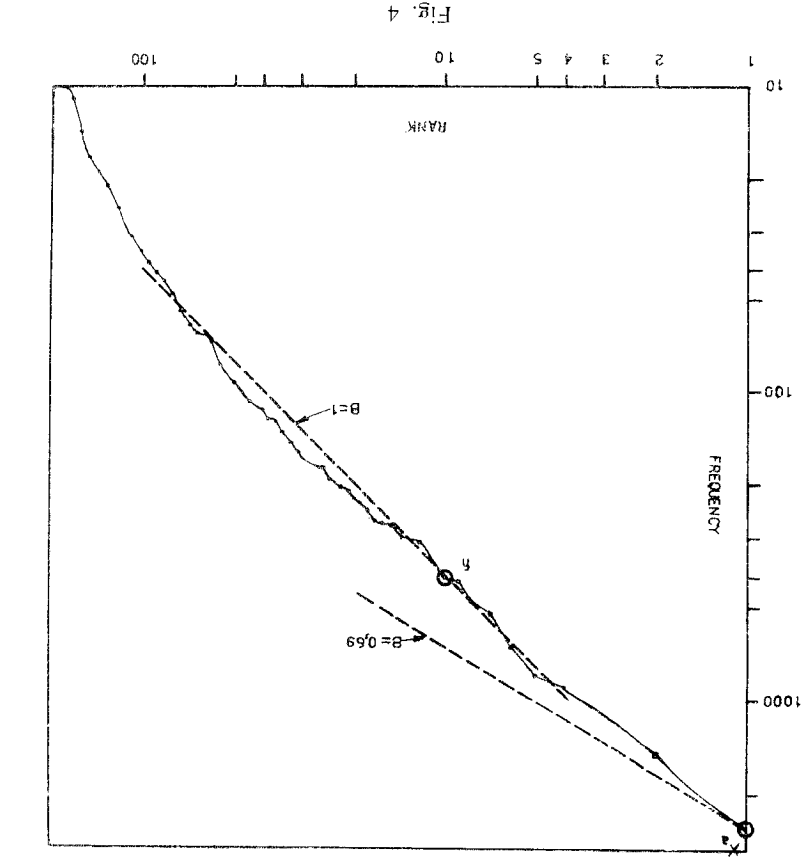


Fig. 4

is 0.48. Since the experimental value of x_n is 3.23 bits, (41) requires $0.48(m - 3.23) = \sigma^2$

A second relation between m and σ is (30). For N is known and the experimental value of x is 10.9 bits. The solution of these equations is $m = 11.12$ bits and $\sigma = 2.8$ bits. The correction factor in (25) is

$$1/a = \Phi \left(\frac{x_n - m}{\sigma} \right) = 0.455$$

and this transforms curve B into curve B' as shown on the right hand side of fig. 6. The cross corresponds to the computed value

of m . This shows that the truncation effect completely accounts for the curvature at the low frequency end.

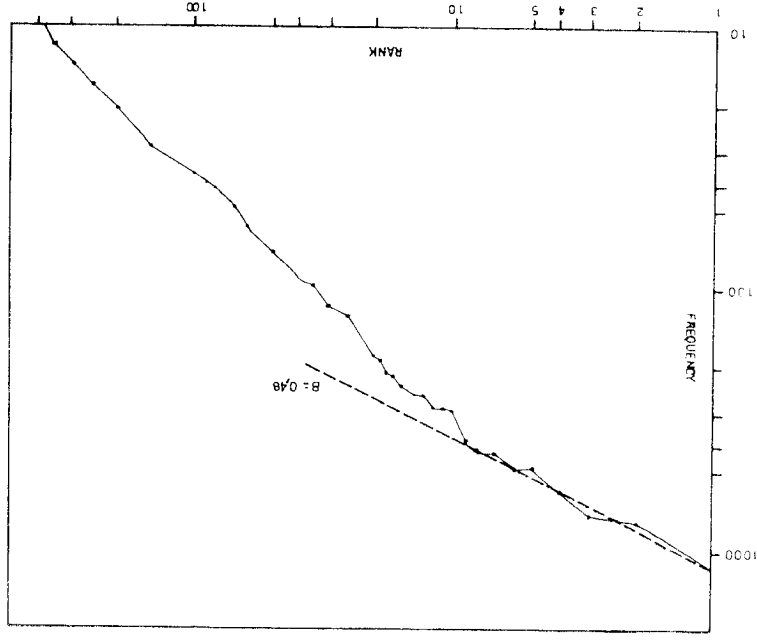


Fig. 5

The various examples of fig. 3 are represented in the left hand side of fig. 6 with an horizontal shift identical to the one discussed in passing from fig. 1 to fig. 2. The smoothed curves are still called A, B, ... E, but the discrete steps have been omitted. It is apparent that all curves will become practically identical after correction for the truncation effect. This correction does not alter the slope of the linear part, and it is already obvious in fig. 3 that all distributions have practically the same slope. The common value seems to be 2.8 bits, which is the double of the value found for phonemes. The straight line in fig. 6 is the theoretical characteristic corresponding to $\sigma = 2.8$ bits.

If one accepts the normal distribution as the general law for words, the fact that Mandelbrot's or Zipf's laws are often satisfactorily confirmed would simply result from the enormous extension of the vocabularies combined with the limitation of many statistics well below the mean rank : for m large, B, as given (41), is constant

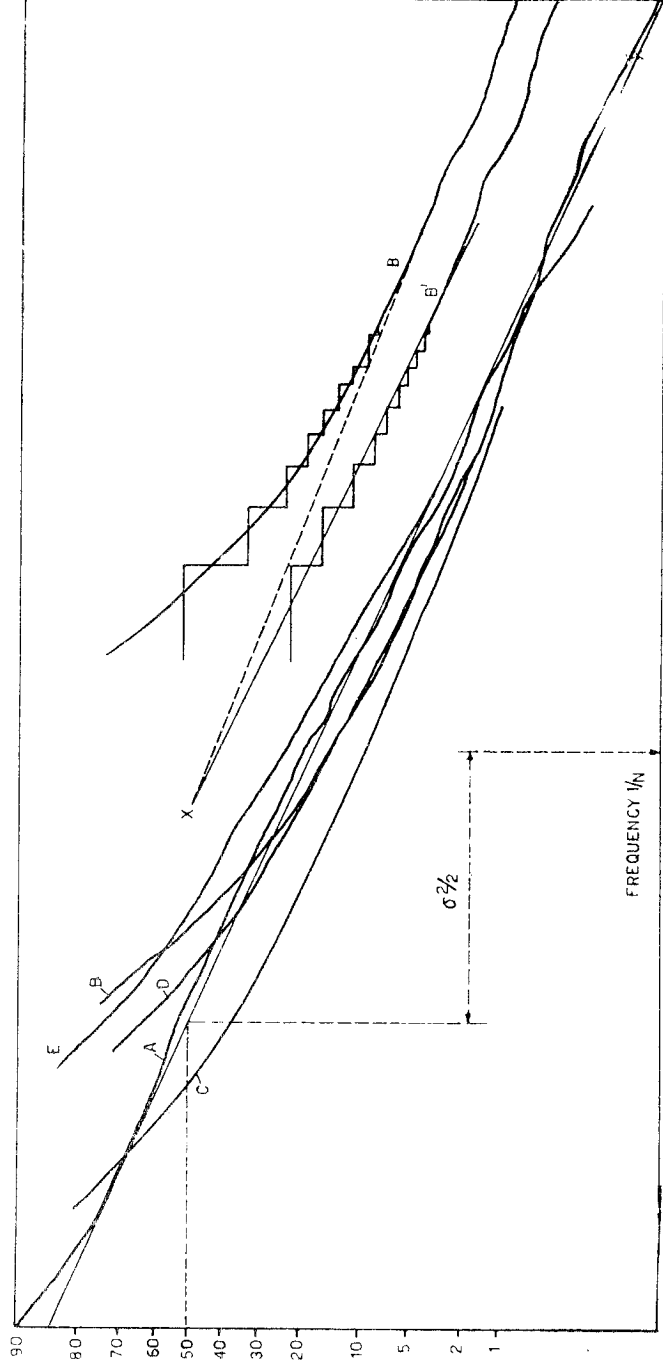


Fig. 6

would also explain the insistence of several authors in counting all inflected forms as distinct.

A number of theoretical models have been proposed by Mandel-rot to account for his law (7). The model based on the weakest hypothesis assumes that texts are separated into words by a randomly distributed „space” symbol. No special theory is needed to explain a normal distribution, but it must be remarked that randomness has been shifted from the text to the vocabulary. For the high frequency tail of the distribution, where the saturation effect due to the finite extent of the vocabulary is still negligible, both explanations are equivalent, since the same stochastic model can be interpreted as yielding the text by a dilution of the dictionary, or the dictionary by a concentration of the text. But, for finite vocabularies a difference arises, because texts remain potentially infinite by hypothesis.

Finally, the constancy of the standard deviation for phoneme distributions on one hand, and for word distributions on the other, would suggest some common discrete substructure for both linguistic levels, but with a double number of degrees of freedom in the latter case.

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L'énergie potentielle complémentaire dans les problèmes dynamiques. - Un principe de variation des accélérations

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SOMMAIRE

Tout comme le principe de variation des déplacements en élasticité, le principe de Hamilton peut être transformé en un principe canonique autorisant des approximations indépendantes sur les déplacements et sur les forces de liaison.

On en déduit un principe dual à celui de Hamilton lorsque les équations d'équilibre dynamique sont exactement satisfaites. Pour des liaisons indépendantes du temps et une énergie cinétique fonction des seules vitesses ce principe dual prend la forme plus élégante d'un principe de variation des accélérations.

Ces idées sont d'abord exposées pour les systèmes à nombre fini de coordonnées lagrangiennes et appliquées, à titre d'exemple, à la recherche de valeurs approchées à la période d'oscillation d'un pendule. Elles sont ensuite étendues au cas de l'élastodynamique linéaire et appliquées à un cas simple de vibrations d'une poutre hyperstatique.

1. GÉNÉRALISATION DU PRINCIPLE DE HAMILTON

Soient $q_r(t) = 1, 2, \dots, n$ des coordonnées généralisées permettant de décrire la configuration d'un système holonome à n degrés de liberté. Le principe de Hamilton (1)

$$(1) \quad \delta \int_{t_2}^{t_1} \mathbb{T}(q_r, \dot{q}_r, t) - V(q_r) dt = 0$$

$$(2) \quad \delta q_r = 0 \text{ pour } t = t_1 \text{ et } t = t_2$$

(7) The publication of B. MANDELROT, « Linguistique statistique macros-copique » in *Logique, langage et théorie de l'information*, Presses Univ. Fr., Paris, 1957, gives a non-mathematical account of his theory, and a bibliography of the subject. See also the more recent contribution of B. MANDELROT, in *Info and Control* vol. 2 pp. 90-99; april 1959.